

# MICROSTRUCTURE OF THE FINANCIAL MARKET

## ILLIQUIDITY EFFECTS IN THE RUSSIAN STOCK MARKET

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# MOTIVATION

- The paradigm of market microstructure invariance proposed by A. Kyle and A. Obizhaeva (2016) allows researchers to better understand the link between market microstructure effects and asset pricing.
- The theory suggests that using an asset-specific illiquidity measure offers several advantages compared to using alternative low-frequency proxies (Amihud's index).
- There is an ambiguity about existence of the illiquidity premium in some emerging markets.

# RESEARCH QUESTIONS

- Is illiquidity a determinant of asset returns in the Russian stock market?
- How large are biases in the estimates of the illiquidity premium?
- What are the potential profits of trading on a measure of stock illiquidity?
- Do the Fama-French and momentum factors fully explain illiquidity premium?

# MARKET MICROSTRUCTURE

“Market microstructure economics focuses on how prices adjust to new information and how the trading mechanism affects asset prices.” (J. R. Russell & R. F. Engle, 2010)

## Topics:

- Price discovery
- Market transparency
- Transaction costs
- Liquidity
- High-frequency trading
- Financial stability

## OVERVIEW

- Securities are different with respect to returns volatility, bid-ask spread, order size, trading volume, etc.
- “Business time” measures the speed at which new bets (sequences of orders of the same direction) arrive.
- **The main idea of the MMI framework:** the distribution of bets and transactions costs are invariant across assets and across time when measured per unit of business time.
- Strong empirical evidence: portfolio transitions by U.S. institutional investors, Thomson Reuters data on financial news articles, intraday patterns for S&P500 E-mini Futures, U.S. stock market trades, etc.

## INVARIANCE-IMPLIED ILLIQUIDITY MEASURE

$$\frac{1}{L_{it}} = \left( \frac{C \cdot \sigma_{it}^2}{m^2 \cdot P_{it} \cdot V_{it}} \right)^{1/3}$$

Notation:

- $C$  - Ruble costs of executing a bet (e.g., 1700 rubles per bet)
- $\sigma_{it}^2$  - Return variance (unitless per day, e.g., 0.06/day)
- $m^2$  - Moment ratio for the distribution of bet sizes (unitless)
- $P_{it}$  - Price (rubles per stock, e.g., 100 rubles/share)
- $V_{it}$  - Volume (stocks per day, e.g., two million stocks per day)

Kyle and Obizhaeva (2016) calibrate  $C = \$2000$  and  $m^2 = 0.25$  for the U.S. stock market using a large sample of portfolio transitions (proxies for bets) over the period 2001–2005.

# COMPARISON BETWEEN METRICS

Amihud's illiquidity measure (a proxy for the response of price to order flow) is similar to

$$\frac{1}{L_{it}^{Amihud}} = \frac{|\sigma_{it}|}{P_{it} \cdot V_{it}} \cdot 10^6.$$

- Kyle and Obizhaeva (2016): the Amihud's ratio implies that the most actively traded and least actively traded assets have the same number of bets per day.
- The Amihud metric does not satisfy **Time Clock Irrelevance**. By contrast,  $1/L_{it}$  remains the same regardless of whether a researcher measures market characteristics using different time horizons. Moreover,  $1/L_{it}$  also satisfies **Modigliani-Miller equivalence**: if each share is bundled with the same amount of cash (levered down by a factor of 2), then  $P_{it} \rightarrow 2 \cdot P_{it}$ ;  $V_{it} \rightarrow V_{it}$ ;  $\sigma_{it} \rightarrow \frac{\sigma_{it}}{2}$ , and  $L_{it}$  doubles as well.

# WHAT $1/L$ MEASURES

Kyle and Obizhaeva (2018):

- The percentage bid-ask spread may be proportional to  $1/L$ ;
- The ruble-weighted average market impact cost of executing bets of the size they are actually executed in the market, expressed as a fraction of the ruble value traded, is exactly equal to  $1/L$ ;
- Market resiliency (the rate at which past pricing errors decay) may be proportional to  $\sigma^2 \cdot L^2$ , etc.



## LITERATURE REVIEW

- **Positive cross-sectional and/or time-series relationship between illiquidity and asset returns:**

*The U.S. stock market:* Amihud and Mendelson (1986), Brennan and Subrahmanyam (1996), Chordia et al. (2001), Pastor and Stambaugh (2003), Acharya and Pedersen (2005), Huh (2014)

*International developed markets:* Lee (2011), Amihud (2015)

*Emerging markets:* Bekaert et al. (2007), Amihud et al. (2015)

- **No illiquidity premium:**

*Emerging markets:* Hearn and Piesse (2008), Hearn et al. (2010), Stereńczak et al. (2020)

- **Positive liquidity premium:**

*Emerging markets:* Batten and Vo (2014), Phong (2016)

Borisenko & Gelman (2012), Mikova (2013):

Liquidity is one of the determinants of asset pricing in the Russian stock market (proxies: trading volume, bid-ask spread, Amihud's ratio)

## DATA

- Thomson Reuters Eikon provides trade-level data (daily opening and closing prices, daily volume) and information about market capitalization and book-to-market ratios.  
Period: January 2010 - December 2019.  
Coverage: 234 ordinary and preferred stocks traded on the Moscow Exchange (before December 2011: MICEX).
- The Moscow Exchange provides historical market data (all trades and best orders) throughout January 2014 - July 2018.

## LIQUIDITY ESTIMATES

The data used for the analysis does not enable us to accurately estimate the  $m^2$  and  $Cq$  parameters for the Russian stock market because they do not contain information on bets.

Assumptions:

1.  $m_{RU}^2 \approx m_{US}^2 = 0.25$
2. Kyle and Obizhaeva (2017):  $C = c \cdot w/b$ , where  $b$  - a finance professional's productivity measured as the number of bets processed per day,  $w$  - the nominal wage for finance professionals per day.  $C$ ,  $w$ , and  $b$  may vary across countries, but  $c$  is a dimensionless constant. We use occupational employment statistics and data on the total number of shares traded.

The final result:  $C_{RU} = C_{US} \cdot \frac{b_{US}}{w_{US}} \cdot \frac{w_{RU}}{b_{RU}} \approx 1700$  rubles. This allows us to redefine the Kyle and Obizhaeva illiquidity measure:

$$1/L_{it} = 30 \cdot \left( \frac{\sigma_{it}^2}{P_{it} \cdot V_{it}} \right)^{1/3}.$$

## ASSET-PRICING TESTS

We exclude firms with negative book values and set  $BM$ ,  $1/L$ ,  $1/L^{Amihud}$  values outside the 0.005 and 0.995 fractiles equal to these fractiles, respectively. For each month in the period February 2012 - December 2019, we run the following cross-sectional regression:

$$R_{it} = \alpha_{0t} + \alpha_{1t}\hat{\beta}_p + \alpha_{2t}SIZE_{it-1} + \alpha_{3t}BM_{it-1} + \alpha_{4t}\Lambda_{it-1} + \epsilon_{it},$$

where  $R_{it}$  is the excess return of stock  $i$  in month  $t$ ,  $\hat{\beta}_p$  - portfolio betas estimated from the full period using 10 portfolios (preranking portfolio betas are estimated for individual stocks using monthly returns from 24 previous months),  $SIZE_{it-1}$  and  $BM_{it-1}$  - logarithms of market capitalization and book-to-market ratios, calculated at the end of the month  $t - 1$ ,  $\Lambda_{it-1}$  is the mean value of  $1/L_{it-1}$  or  $1/L_{it-1}^{Amihud}$  for the month  $t - 1$ . The coefficients from the cross-sectional regressions are averaged over time, using the Fama-MacBeth (1973) methodology.

## DESCRIPTIVE STATISTICS

Variable	Mean	Median	Min	Max
<i>Return</i>	-7.13	-8.19	-210.66	255.88
$\hat{\beta}_p$	0.49	0.51	-0.21	1.01
<i>Size</i>	23.43	23.45	16.50	29.45
<i>BM</i>	0.23	0.29	-3.33	2.73
$1/L \cdot 10^{-4}$	291	184	7.99	1932
$1/L^{Amihud} \cdot 10^{-4}$	406	28	0.003	880

**TABLE:** This table displays means, medians, minimum values, and maximum values on the variables included in the asset-pricing tests. All statistics are calculated from the full sample. *Return* - percentage monthly return in excess of the Russia's 1 month government bond yield;  $\hat{\beta}_p$  - portfolio betas estimated from the full period using 10 portfolios; *Size* - logarithm of market capitalization; *BM* - logarithm of the book value of equity divided by the market value of equity;  $1/L$  and  $1/L^{Amihud}$  are the Kyle-Obizhaeva measure and the Amihud illiquidity ratio, respectively.


## CORRELATIONS

	$\hat{\beta}_p$	<i>SIZE</i>	<i>BM</i>	<i>1/L</i>	<i>1/L</i> <sup><i>Amihud</i></sup>
<i>Return</i>	-0.053	-0.066	-0.049	0.183	0.134
$\hat{\beta}_p$		0.100	0.024	-0.148	-0.105
<i>SIZE</i>			-0.233	-0.668	-0.404
<i>BM</i>				0.091	0.019
<i>1/L</i>					0.852

**TABLE:** This table displays the time-series means of monthly bivariate correlations of the variables in the asset-pricing tests.

## FAMA-MACBETH REGRESSION ESTIMATES

	2012-2019		2012-2019 (Excl. January)	
	$\Lambda = 1/L$	$\Lambda = 1/L^{Amihud}$	$\Lambda = 1/L$	$\Lambda = 1/L^{Amihud}$
<i>Constant</i>	-12.515 (-4.543)	0.312 (0.139)	-11.172 (-4.442)	1.451 (0.611)
$\hat{\beta}_p$	-1.657 (-2.327)	-1.949 (-2.775)	-1.886 (-2.568)	-2.2072 (-3.054)
<i>SIZE</i>	0.154 (1.454)	-0.304 (-3.157)	0.099 (0.991)	-0.351 (-3.357)
<i>BM</i>	-0.226 (-0.879)	-0.259 (-1.049)	-0.400 (-1.232)	-0.377 (-1.416)
<i>1/L</i>	<b>90.815</b> (8.951)		<b>88.895</b> (8.190)	
<i>1/L<sup>Amihud</sup></i>		<b>20.411</b> (6.800)		<b>19.145</b> (6.147)
R-squared (overall)	0.0236	0.0171	0.0251	0.0186
Stocks	203	203	203	203
Time periods	95	95	88	88
No. Obs.	12065	12065	11221	11221

T-values (in parentheses) are adjusted with the Newey-West procedure with 3 lags. 

## DISCUSSION

- Interpretation of the estimate of the  $1/L$  coefficient: a difference of 1 percentage point in the average price impact cost (expressed in percentage terms) between two stocks translates into a difference in required return of **0.908%** per month.
- Interpretation of the estimate of the  $1/L^{Amihud}$  coefficient: a difference of 1 percentage point in the time-series average of the monthly ratios (as multiples of  $10^{-4}$ ) of percentage returns in absolute magnitude to ruble volume between two stocks translates into a difference in required return of **0.204%** per month. Implicit assumption: the number of bets across assets is constant.
- The coefficients on  $1/L$  and  $1/L^{Amihud}$  remain positive and significant for each of the two subperiods (2012-2015 and 2016-2019).
- Some coefficients on other variables are inconsistent with the standard asset-pricing theory. Errors-in-variables bias because of estimation errors in the betas from the first stage?  
T.Teplova & V.Rodina suggest using an approach to calculating beta with liquidity adjustment: a shift from historical to predicted beta.



## FURTHER RESEARCH

- The cross-sectional/time-series illiquidity effects in the U.S. stock market with the Kyle and Obizhaeva measure as a proxy.
- The time-series regression approach of Fama and French (1993) for studying illiquidity factors in the Russian stock market.
- The trading cost model of Kyle and Obizhaeva (2016) for investigating the “true” profitability of strategies that require a regular rebalancing of the portfolio (e.g., momentum).

# CONCLUSIONS

- Our tests support the predictions of the cross-sectional relationship between stock return and stock illiquidity in the Russian equity market.
- Of particular importance is conducting several robustness checks: alternative estimation procedures (e.g., GMM), different approaches to calculating beta (e.g., with liquidity adjustment) and return variance (e.g., the log-ARIMA model).