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ADVANCED ECONOMETRICS

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Univariate Time Series Models

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General ARMA processes

$$y_t = Y_t - \mu,$$

AR(p) process:

$$y_t = \theta_1 y_{t-1} + \theta_2 y_{t-2} + \dots + \theta_p y_{t-p} + \varepsilon_t,$$

MA(q) process:

$$y_t = \varepsilon_t + \alpha_1 \varepsilon_{t-1} + \dots + \alpha_q \varepsilon_{t-q},$$

ARMA(p,q) model:

$$y_t = \theta_1 y_{t-1} + \theta_2 y_{t-2} + \dots + \theta_p y_{t-p} + \varepsilon_t + \alpha_1 \varepsilon_{t-1} + \dots + \alpha_q \varepsilon_{t-q}$$

Lag operator :

$$L(Y_t) = Y_{t-1}$$

$$L^s(Y_t) = Y_{t-s}$$

$$AR(1): y_t = \theta y_{t-1} + \varepsilon_t \Leftrightarrow (1 - \theta L)y_t = \varepsilon_t$$

$$AR(p): y_t = \theta_1 y_{t-1} + \theta_2 y_{t-2} + \dots + \theta_p y_{t-p} + \varepsilon_t \Leftrightarrow$$

$$\theta(L)y_t = \varepsilon_t,$$

$$\theta(L) = 1 - \theta_1 L - \theta_2 L^2 - \dots - \theta_p L^p$$

Invertability of Lag polynomials

$$AR(1): y_t = \theta y_{t-1} + \varepsilon_t \Leftrightarrow (1 - \theta L)y_t = \varepsilon_t,$$

$$(1 - \theta L)^{-1}(1 - \theta L)y_t = (1 - \theta L)^{-1}\varepsilon_t,$$

$$(1 - \theta L)^{-1} = \sum_{j=0}^{\infty} \theta^j L^j \quad \text{if } |\theta| < 1,$$

($|\theta| < 1$ is the condition for the invertibility),

$$y_t = (1 - \theta L)^{-1}\varepsilon_t,$$

$$y_t = \sum_{j=0}^{\infty} \theta^j \varepsilon_{t-j}.$$

The MA representation is convenient to determine variance and covariance.

$$AR(2): y_t = \theta_1 y_{t-1} + \theta_2 y_{t-2} + \varepsilon_t \Leftrightarrow$$

$$(1 - \theta_1 L - \theta_2 L^2) y_t = \varepsilon_t,$$

$$(1 - \varphi_1 L)(1 - \varphi_2 L) y_t = \varepsilon_t,$$

The condition for invertability of $(1 - \varphi_1 L)(1 - \varphi_2 L)$:

both $(1 - \varphi_1 L)$ and $(1 - \varphi_2 L)$ must be invertible.

Thus, the requirement for invertability:

$$|\varphi_1| < 1, \quad |\varphi_2| < 1.$$

$$AR(2): y_t = \theta_1 y_{t-1} + \theta_2 y_{t-2} + \varepsilon_t \Leftrightarrow$$

$$(1 - \varphi_1 L)(1 - \varphi_2 L)y_t = \varepsilon_t,$$

The characteristic equation:

$$(1 - \varphi_1 z)(1 - \varphi_2 z) = 0.$$

Conditions for invertability

$|\varphi_1| < 1, |\varphi_2| < 1$ correspond to

$$|z_1| > 1, |z_2| > 1.$$

*The AR polynomial is invertable
if and only if the process is stationary.*

Invertability of MA polynomial

$$MA(1): y_t = \varepsilon_t + \alpha\varepsilon_{t-1} \Leftrightarrow$$

$$y_t = (1 + \alpha L)\varepsilon_t,$$

$$(1 + \alpha L)^{-1} = \sum_{j=0}^{\infty} (-\alpha)^j L^j,$$

$$y_t = \alpha \sum_{j=0}^{\infty} (-\alpha)^j y_{t-j-1} + \varepsilon_t$$

$$MA(q): y_t = \alpha(L)\varepsilon_t.$$

*A necessary condition for the $AR(\infty)$ representation:
invertability of $\alpha(L)$.*

Stationarity and unit root

ARMA (p, q) model :

$$\theta(L)y_t = \alpha(L)\varepsilon_t,$$

corresponds to a stationary process if and only if the solutions of equations

$$\theta(z) = 0$$

satisfy conditions : $|z_j| > 1 \quad \forall j = 1, \dots, p.$

Special case : $\theta(L) = \theta^(L)(1-L),$*

$\theta^(L)$ is an invertable polynomial .*

$$\theta^*(L)(1-L)y_t = \alpha(L)\varepsilon_t \Rightarrow$$

$\theta^(L)\Delta y_t = \alpha(L)\varepsilon_t$ is a stationary process .*

Stationarity and unit root

A stationary series is denoted by $I(0)$.

A series which become stationary after first differencing is said to be integrated of order one, denoted $I(1)$.

A series which become stationary only after second differencing is said to be integrated of order two, denoted $I(2)$, etc.

$$AR(1) : Y_t = \delta + \theta Y_{t-1} + \varepsilon_t$$

$$H_0 : \theta = 1 \quad (\text{a unit root})$$

$$H_1 : |\theta| < 1 \quad (\text{stationarity})$$

Dickey–Fuller test :

$$DF = \frac{\hat{\theta} - 1}{se(\hat{\theta})}$$

Testing for Unit Roots

More convenient procedure

$$\Delta Y_t = \delta + (\theta - 1)Y_{t-1} + \varepsilon_t$$

$\theta = 1$ corresponds

$\Delta Y_t = \delta + \varepsilon_t$ – random walk with drift.

Y_t – difference stationary.

Testing for Unit Roots

It is possible that nonstationary is caused by the presence of the deterministic time trend :

$$Y_t = \delta + \theta Y_{t-1} + \gamma t + \varepsilon_t$$

with $|\theta| < 1$ and $\gamma \neq 0$.

Y_t – trend stationary.

Usual DF test :

$$\Delta Y_t = \delta + (\theta - 1)Y_{t-1} + \gamma t + \varepsilon_t,$$

different critical values of DF statistics : $\tau_0, \tau_\mu, \tau_\tau$.

Kwiatkowski, D., Phillips, P. C., Schmidt, P., & Shin, Y. (1992). Testing the null hypothesis of stationarity against the alternative of a unit root: How sure are we that economic time series have a unit root?. *Journal of econometrics*, 54(1), 159-178.

$$y_t = \beta t + r_t + \varepsilon_t$$

$$r_t = r_{t-1} + u_t$$

$$H_0 : \sigma_u^2 = 0 \text{ (stationarity)}$$

$$H_1 : \sigma_u^2 \neq 0 \text{ (non stationarity),}$$

$H_0 : \sigma_u^2 = 0$ (*stationarity*)

$H_1 : \sigma_u^2 \neq 0$ (*non stationarity*),

OLS, residuals e_t , $S_t = \sum_{s=1}^T e_s$,

$$KPSS = \sum_{s=1}^T S_t^2 / \hat{\sigma}^2.$$

The asymptotic distribution is non-standard,

and Kwiatkowski et al. report $KPSS_{0.05}^{cr} = 0.146$.

The same algorithm for model without trend,

$$KPSS_{0.05}^{cr} = 0.463.$$

Testing for Unit Roots in Higher Order AR Models

AR(2) process : $Y_t = \delta + \theta_1 Y_{t-1} + \theta_2 Y_{t-2} + \varepsilon_t,$

$\Delta Y_t = \delta + (\theta_1 + \theta_2 - 1)Y_{t-1} - \theta_2 \Delta Y_{t-1} + \varepsilon_t.$

$H_0 : \pi \equiv \theta_1 + \theta_2 - 1 = 0$ (*unit root*)

AR(p) process : $Y_t = \delta + \theta_1 Y_{t-1} + \dots + \theta_p Y_{t-p} + \varepsilon_t,$

$\Delta Y_t = \delta + \pi Y_{t-1} + c_1 \Delta Y_{t-1} + \dots + c_{p-1} \Delta Y_{t-p+1} + \eta + \varepsilon_t.$

Augmented Dickey – Fuller test

$H_0 : \pi \equiv \theta_1 + \dots + \theta_p - 1 = 0$ (*unit root*).

Test statistics $\tau_0, \tau_\mu, \tau_\tau.$

Nonparametric Philips and Perron test.

Dolado, Jenkinson, Sosvilla-Rivero testing for Unit Roots

Step1. $\Delta Y_t = \delta + \gamma t + \pi Y_{t-1} + c_1 \Delta Y_{t-1} + \dots + c_{p-1} \Delta Y_{t-p+1} + \varepsilon_t.$

$H_0 : \pi = 0$ (unit root, DS),

$H_1 : \pi < 0$ (TS).

If $t < \tau_\tau$ (H_0 is rejected) \Rightarrow TS, no unit root.

If $t > \tau_\tau$ (H_0 is not rejected) go to step 2.

Step2. $\Delta Y_t = \delta + \gamma t + c_1 \Delta Y_{t-1} + \dots + c_{p-1} \Delta Y_{t-p+1} + \varepsilon_t.$

$H_0 : \gamma = 0,$

$H_1 : \gamma \neq 0.$

If H_0 is rejected, stop.

If H_0 is not rejected, go to step 3.

Step3. $\Delta Y_t = \delta + \pi Y_{t-1} + c_1 \Delta Y_{t-1} + \dots + c_{p-1} \Delta Y_{t-p+1} + \varepsilon_t.$

$H_0 : \pi = 0$ (unit root, DS),

$H_1 : \pi < 0$ (const).

If $t < \tau_\mu$ (H_0 is rejected) \Rightarrow no unit root, no time trend.

If $t > \tau_\mu$ (H_0 is not rejected) go to step 4.

Step4. $\Delta Y_t = \delta + c_1 \Delta Y_{t-1} + \dots + c_{p-1} \Delta Y_{t-p+1} + \varepsilon_t.$

$H_0 : \delta = 0,$

$H_1 : \delta \neq 0.$

If H_0 is rejected, stop.

If H_0 is not rejected, go to step 5.

Dolado, Jenkinson, Sosvilla-Rivero testing for Unit Roots

Step5. $\Delta Y_t = \pi Y_{t-1} + c_1 \Delta Y_{t-1} + \dots + c_{p-1} \Delta Y_{t-p+1} + \varepsilon_t.$

$H_0 : \pi = 0$ (*unit root, DS*),

$H_1 : \pi < 0$ (*stationarity*).

If $t < \tau_0$ (H_0 is rejected) \Rightarrow *no unit root.*

If $t > \tau_0$ (H_0 is not rejected) \Rightarrow *unit root..*

Autocorrelation function (ACF):

$$\rho_k = \frac{\text{cov}\{Y_t, Y_{t-k}\}}{\text{var}\{Y_t\}} = \frac{\gamma_k}{\gamma_0},$$

Partial Autocorrelation function (PACF).

The k -th order sample partial autocorrelation coefficient is the estimate for θ_k in an $AR(k)$ model.

AR(2) process: $Y_t = \delta + \theta_1 Y_{t-1} + \theta_2 Y_{t-2} + \varepsilon_t,$

Estimating $Y_t = \delta + \theta_1 Y_{t-1} + \varepsilon_t$ gives $\hat{\theta}_{11},$

estimating $Y_t = \delta + \theta_1 Y_{t-1} + \theta_2 Y_{t-2} + \varepsilon_t$ gives $\hat{\theta}_{22}.$

Partial Autocorrelation Function

AR(p) process: PACF = 0 if $k > p$.

*MA(q) process = AR(∞) if MA polynomial is invertible,
PACF coefficients will damp toward zero.*

AR(p) process:

- 1) An ACF is infinite in extend (it tails off),
- 2) A PACF is close to zero for lags larger than p

MA(q) process:

- 1) An ACF is close to zero for lags larger than q ,
- 2) A PACF is infinite in extend (it tails off),

ARMA(p, q), ARMA(p + 1, q), ARMA(p, q + 1) or ???

Idea: residuals of an adequate model should be approximately white noise.

The significance of residual autocorrelation is often checked by comparing with $\pm 2/\sqrt{T}$

$$H_0 : ARMA(p, q)$$

Ljung–Box portmanteau test statistics:

$$Q_K = T(T + 2) \sum_{k=1}^K \frac{1}{T - k} r_k^2,$$

r_k^2 are the estimated autocorrelation coefficient
of the residuals $\hat{\varepsilon}_t$

and K is a number chosen by the researcher.

$$H_0 : ARMA(p, q) \Rightarrow Q_K \sim \chi^2(K - p - q - 1).$$

Criteria for model selection

Akaike's Information Criterion (AIC):

$$AIC = \ln \hat{\sigma}^2 + 2 \frac{p+q+1}{T}.$$

Schwarz's Bayesian Information Criterion (BIC):

$$BIC = \ln \hat{\sigma}^2 + \frac{p+q+1}{T} \ln T.$$

ARIMA(p, d, q)

- 1) d – order of difference (after UR testing).
- 2) Analysis of ACF, PACF $\rightarrow p, q$
- 3) AIC, BIC
- 4) Estimation of parameters
- 5) AR – OLS
- 6) MA – ML
- 7) ARMA - ML