$$Y^* = \beta_1 + \beta_2 X + u$$

Sometimes the dependent variable in a regression model is subject to a lower limit or an upper limit, or both. Suppose that in the absence of any constraints, *Y* is related to *X* by the model shown.



For example, suppose that the true relationship is as shown, and that, given a set of random numbers for the disturbance term drawn from a normal distribution with mean 0 and standard deviation 10, we have a sample of observations as shown.

$$Y^* = \beta_1 + \beta_2 X + u$$
$$Y = Y^* \text{ if } Y^* > 0$$
$$Y = 0 \quad \text{if } Y^* \le 0$$

However, suppose that the dependent variable is subject to a lower bound, in this case 0. Then Y will be as given by the model if  $Y^* > 0$ , and it will be 0 if  $Y^* = 0$  or if  $Y^* < 0$ .



For example, suppose that we have a labor supply model with Y hours of labor supplied per week as a function of X, the wage that is offered. It is not possible to supply a negative number of hours.



Those individuals with negative Y\* will simply not work. For them, the actual Y is 0.



What would happen if we ran an OLS regression anyway? Obviously, in this case the slope coefficient would be biased downwards.

$$Y^* = \beta_1 + \beta_2 X + u$$

$$Y = Y^*$$
 if  $Y^* > 0$ 

observation dropped if  $Y^* \leq 0$ 

It would be natural to suppose that the problem could be avoided by dropping the constrained observations. Unfortunately, this does not work.



Here again is the sample as it would be if Y were not constrained.



Here is the sample with the constrained observations dropped.



An OLS regression again yields a downwards-biased estimate of the slope coefficient and an upwards-biased estimate of the intercept. We will investigate the reason for this.



Look at the two observations highlighted. For such low values of *X*, most of the observations are constrained. The reason that these two observations appear in the sample is that their disturbance terms happen to be positive and large.



In general, for an observation to appear in the sample,  $Y^*$  must be positive, and this requires that u > 40 - 1.2X.



If X is equal to 10, u must be greater than 28 if the observation is to appear in the sample.



If *X* is equal to 20, *u* must be greater than 16.



If *X* is equal to 30, *u* must be greater than 4.



If X is equal to 40, the observation will appear in the sample for any positive value of u and even some negative ones. The condition is that u must be greater than -8.



If *X* is equal to 50, *u* must be greater than –20.



If *X* is equal to 60, *u* must be greater than –32. A value of less than –32 is very unlikely, so in this part of the sample virtually every observation will appear.



We will now show that, for observations that appear in the sample, there is a negative correlation between *X* and *u*.



When X is equal to 10, *u* must be greater than 28. The expected value of *u* for observations appearing in the sample is its expected value in the tail to the right of the red line. It turns out to be 31.0.



When X is equal to 20, *u* must be greater than 16. The expected value of *u*, for observations that appear in the sample, is the expected value in the shaded area. This is 20.2.



The rest of the distribution is irrelevant because an observation cannot appear in the sample if u < 16.



When X is equal to 30, *u* must be greater than 4. Its expected value, conditional on it being greater than 4, is 10.7.



When X is equal to 40, *u* must be greater than –8. Its expected value, subject to this condition, is 3.7.



When X is equal to 50, u must be greater than -20. It will satisfy this condition nearly all the time and its conditional expected value, 0.6, is hardly any greater than its unconditional expected value, 0.



When X is 60 or higher, the condition will always be satisfied and the observation will always appear in the sample. For X > 60, E(u) is equal to its unconditional value of 0 and so there is no negative correlation between X and u.



The range over which one observes a negative correlation between X and u is approximately 15 to 45. Below 15, an observation is almost certainly going to be constrained and so deleted from the sample.



Above 45, almost all are going to appear in the sample, irrespective of the value of *u*.



The solution to the problem is to have a hybrid model which effectively uses probit analysis to investigate why some observations have positive  $Y^*$  while others do not, and then, for those with  $Y^* > 0$ , regression analysis to quantify the relationship.



The model is fitted using maximum likelihood estimation. We will not be concerned with the technicalities here.



We will use the Consumer Expenditure Survey data set to illustrate the use of tobit analysis. The figure plots annual household expenditure on household equipment, *HEQ*, on total household expenditure, *EXP*, both measured in dollars.

. tab HEQ if HEQ<10

HEQ	Freq.	Percent	Cum.
0	86	 89.58	 89.58
3	1	1.04	90.62
4	2	2.08	92.71
6	1	1.04	93.75
7	1	1.04	94.79
8	5	5.21	100.00
Total	 96	100.00	

For 86 households, *HEQ* was 0. (The tabulation has been confined to small values of *HEQ*. We are only interested in finding out how many actually had HEQ = 0.)

. reg HEQ EXP

Source	SS	df	MS		Number of obs	= 869 - 353 01
Model   Residual   + Total	729289164 1.7866e+09 2.5159e+09	1 7 867 20 	29289164 60635.12  98456.01		F( I, 807) Prob > F R-squared Adj R-squared Root MSE	= 0.33.91 $= 0.0000$ $= 0.2899$ $= 0.2891$ $= 1435.5$
HEQ	Coef.	Std. Err	t. t	P> t	 [95% Conf.	Interval]
EXP   cons	.0471546 -397.2088	.0025065	18.81	3 0.000 1 0.000	.042235 -572.7619	.0520742 -221.6558

Here is a regression using all the observations. We anticipate that the coefficient of *EXP* is biased downwards.

. reg HEQ EXP if HEQ>0

Source	I	SS	df		MS		Number of obs	=	783
	+-						F(1, 781)	=	291.04
Model	I	656349265	1	6563	49265		Prob > F	=	0.0000
Residual	I	1.7613e+09	781	22552	19.19		R-squared	=	0.2715
	+-						Adj R-squared	=	0.2705
Total	I	2.4177e+09	782	30916	56.59		Root MSE	=	1501.7
HEQ	   +-	Coef.	Std.	Err.	t	P> t	[95% Conf.	In	terval]
EXP	т- 	.0467672	.0027	7414	17.060	0.000	.0413859	•	0521485
_cons		-350.1704	101.8	3034	-3.440	0.001	-550.0112	-1	50.3296

Here is an OLS regression with the constrained observations dropped. The estimate of the slope coefficient is almost the same, just a little lower.

. tobit HEQ EXP, 11(0)

chi2(1) = Prob > chi2 = Log Likelihood = -6911.0175 Pseudo R2 =	= 315.41 = 0.0000 = 0.0223
Log Likelihood = -6911.0175 Prob > chi2 = Prob > chi2 = Pseudo R2 =	= 0.0000 = 0.0223
Log Likelihood = -6911.0175 Pseudo R2 =	: 0.0223
	0.0223
HEQ   Coef. Std. Err. t P> t  [95% Conf. In	iterval]
EXP   .0520828 .0027023 19.273 0.000 .0467789 .	0573866
_cons   -661.8156 97.95977 -6.756 0.000 -854.0813 -4	69.5499
Obs. summary: $86$ left-censored observations at HEOC=0	
783 uncensored observations	

Here is the tobit regression. The Stata command is 'tobit', followed by the dependent variable and the explanatory variables, then a comma, then '11' and in parentheses the lower limit.

. tobit HEQ EXP, 11(0)

Tobi	t Estin	nates				Number of ob	s = 869
Log	Likelił	nood = -6911	0175			chi2(1) Prob > chi2 Pseudo R2	$= 315.41 \\ = 0.0000 \\ = 0.0223$
	HEQ	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
	EXP	. 0520828	.0027023	19.273	0.000	.0467789	.0573866
	cons	-661.8156	97.95977	-6.756	0.000	-854.0813	-469.5499
	_se	1521.896	38.6333		(Ancilla	ary parameter)	
Obs.	summaı	cy: 80 783	5 left-censon 8 uncensored	red observa observatio	ations at	HEQ<=0	

If the dependent variable were constrained by an upper limit, we would use 'u1' instead of '11', with the upper limit in parentheses. The method can handle lower limits and upper limits simultaneously.

. tobit HEQ EXP, 11(0)					
HEQ   Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
EXP   .0520828 cons   -661.8156	.0027023 97.95977	19.273 -6.756	0.000 0.000	.0467789 -854.0813	.0573866 -469.5499
_se   1521.896	38.6333		(Ancillary	parameter)	
. reg HEQ EXP					
HEQ   Coef.	Std. Err.	t	₽> t	[95% Conf.	Interval]
EXP   .0471546 cons   -397.2088	.0025065 89.44449	18.813 -4.441	0.000 0.000	.042235 -572.7619	.0520742 -221.6558
. reg HEQ EXP if HEQ>0					
HEQ   Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
EXP   .0467672 _cons   -350.1704	.0027414 101.8034	17.060 -3.440	0.000 0.001	.0413859 -550.0112	.0521485 -150.3296

We see that the coefficient of *EXP* is indeed larger in the tobit analysis, confirming the downwards bias in the OLS estimates. In this case the difference is not very great. That is because only 10 percent of the observations were constrained.

Copyright Christopher Dougherty 2001–2006. This slideshow may be freely copied for personal use.